

imputation procedure is reported in Figure 3b. Our third strategy is to impute the missing data using information on the spatial distribution of ethnic groups from the *Geo-Referencing of Ethnic Groups* (GREG) database (Weidmann, Rod and Cederman, 2010). Like the *Ethnologue*, the GREG database provides a shape file that divides the world’s land into polygons, with each polygon indicating the location of a specific ethnicity. The shortcoming of the GREG database is that ethnic groups are much less finely identified relative to the *Ethnologue* database. The GREG database identifies 1,364 ethnic groups, while the *Ethnologue* identifies 7,612 language groups.¹³ The spatial distribution of historical plough use using this procedure is shown in Figure 3c.

In Figures 4a–4c, we report population weighted country-level averages of historical plough use for each of the three strategies used to address the missing language data. In our analysis, we use the plough variable that was constructed without missing values imputed as our baseline measure. Our results are robust to the use of either variable that imputes the missing language data. This robustness is explained by the high correlation among the three plough measures. At the country-level, the correlation between: (i) our baseline variable and the measure with missing languages imputed using the country’s national language is 0.89; (ii) our baseline measure and the measure imputed using ethnic groups from the GREG database is 0.91; and (iii) the two variables with imputed values is 0.99.¹⁴

4. OLS estimates

Having constructed country- and district-level measures of traditional plough use, we are able to examine the relationship between historical plough use and the role of women in societies today. We begin by examining variation at the country level.

A. Country-level estimates

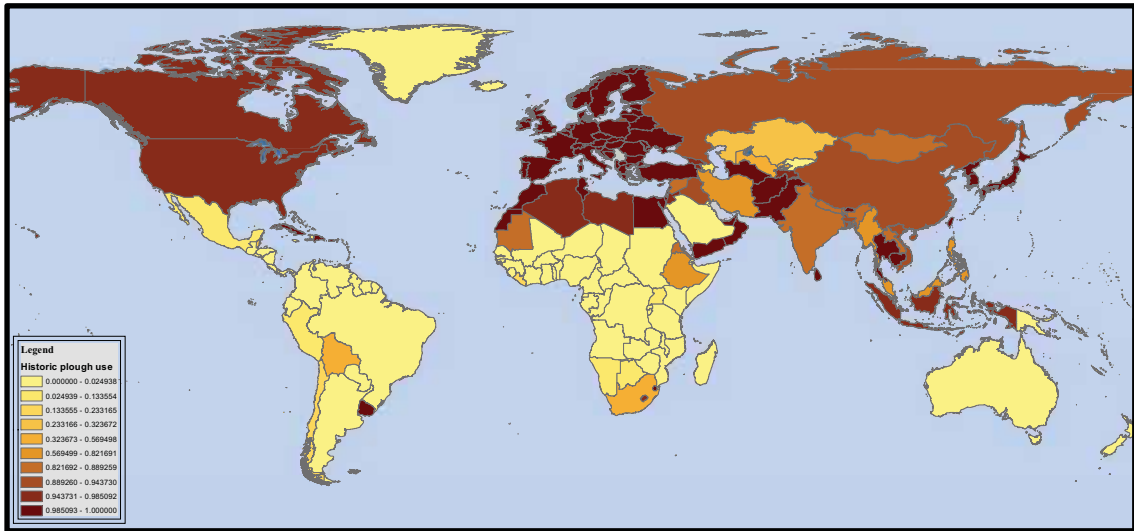
We test Boserup’s hypothesis by estimating the following equation:

$$y_c = \alpha + \beta \text{Plough}_c + \mathbf{X}_c^C \boldsymbol{\Gamma} + \mathbf{X}_c^H \boldsymbol{\Pi} + \varepsilon_c \quad (2)$$

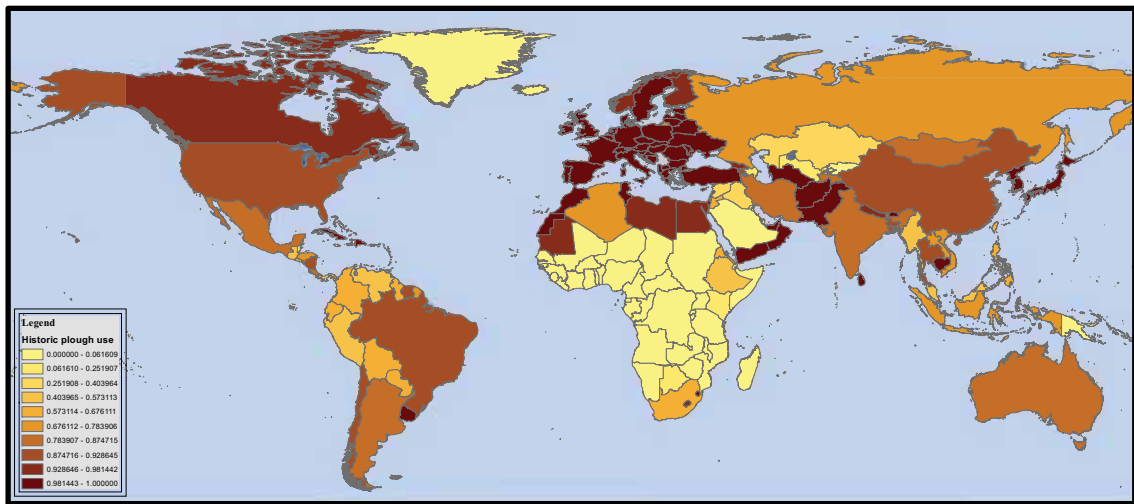
where y is an outcome of interest, c denotes countries, Plough_c is our measure of the historical use of the plough among the ancestors of the citizens in country c , and \mathbf{X}_c^C and \mathbf{X}_c^H are vectors

¹³An alternative strategy is to rely only on the coarser GREG classification and map. Our results are robust to this procedure as well.

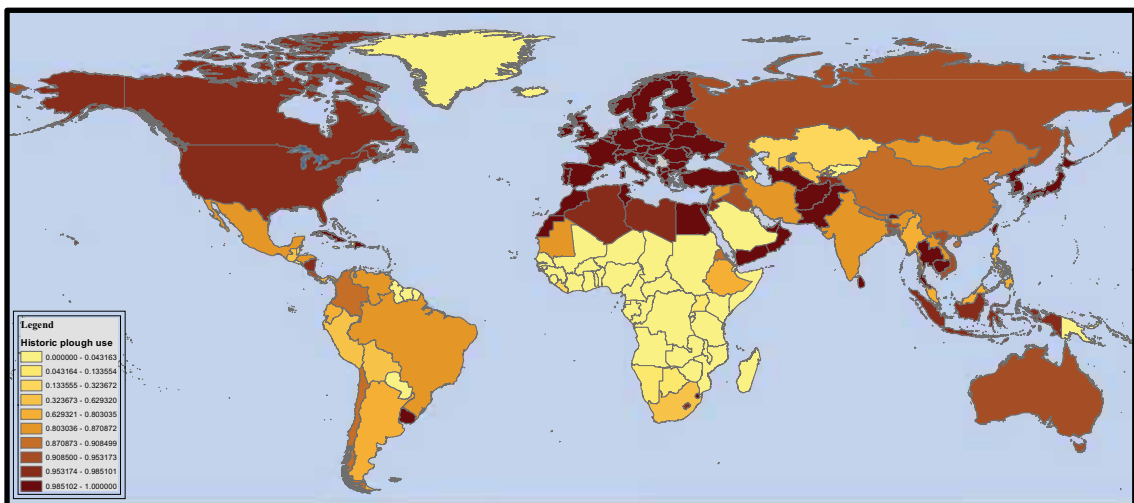
¹⁴Descriptive statistics for the three measures are reported in Appendix Table A1.



(a) Missing language information not imputed



(b) Missing language information imputed using the country's official language



(c) Missing language information imputed using GREG ethnic groups

Figure 4: Average historical plough use among the ancestors of each country

of current controls and historical ethnographic controls, all measured at the country level. X_c^C includes the natural log of a country's real per capita GDP measured in 2000, as well as the variable squared. These controls are important given the well-established non-linear U-shaped relationship between economic development and female labor force participation (Goldin, 1995). We also include an indicator variable that equals one if the country was formerly communist, since these regimes implemented policies to eliminate gender differences in the economy.¹⁵ The historical ethnographic controls X_c^H are based on the ethnographic characteristics used as controls in the ethnicity-level regressions: the presence of domesticated bovine or equine animals, economic development measured by the density of settlement, levels of political authority in the society, agricultural suitability, and the presence of a tropical climate. We construct country-level versions of these variables using the same procedure that is used to construct the historical plough use variable. Thus, the ethnographic controls capture the historical characteristics of a location's ancestors.

Table 2 reports country-level OLS estimates. In columns 1 and 2, the dependent variable is a country's female labor force participation rate (FLFP) in 2000.¹⁶ In columns 3–6, we examine women's participation in more narrowly specified activities outside of the domestic sphere: entrepreneurship (measured by the share of firms with owners or managers that are female) and national politics (measured by the proportion of seats held by women in national parliament).¹⁷ The even numbered columns include controls for continent fixed effects, while the odd numbered columns do not. The estimates show that in countries with a tradition of plough use, women are less likely to participate in the labor market, are less likely to own or manage firms, and are less likely to participate in politics.¹⁸ All coefficients are negative and statistically significant.

The partial correlation plots for historical plough use are show in Figures 5a–5c (for columns

¹⁵Alesina and Fuchs-Schundeln (2007) show how the impact of a communist regime on individual beliefs can be long lasting.

¹⁶FLFP is taken from the World Bank's *World Development Indicators*. The variable is measured as the percentage of women aged 15 to 64 that are in the labor force.

¹⁷The share of a country's firms with some female ownership is measured as the percentage of surveyed firms with a woman among the principal owners. The data are from the World Bank *Enterprise Surveys*. The proportion of seats in national parliament is measured as the percent of parliamentary seats, in a single or lower chamber, held by women. The variable, measured in 2000, is taken from the United Nations' *Women's Indicators and Statistics Database*.

¹⁸Because female participation in national politics may be affected by the type of government, we also control for each country's level of democracy in 2000 when this outcome is examined. Democracy is measured using the 'polity2' measure from the Polity IV database, which takes on integer values and ranges from –10 (high autocratic) to +10 (highly democratic). Many countries have introduced quotas to increase the participation of women in politics. We have checked the robustness of our results excluding countries with gender quotas. For the restricted sample of 86 countries, the estimated coefficient is –4.85 with a standard error of 2.68.

Table 2: Country level OLS estimates.

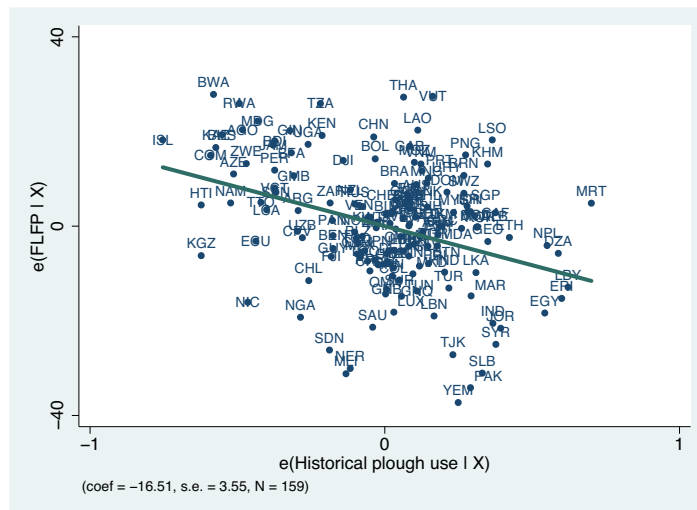
	Dependent variable:							
	Female labor force participation		Share of firms with some female ownership		Females in politics		Average effect size (AES)	
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Historical plough use	-16.506*** (3.547)	-15.417*** (3.561)	-11.052** (4.287)	-11.540** (5.152)	-5.606*** (2.128)	-4.245* (2.218)	-0.849*** (0.140)	-0.796*** (0.137)
<i>Historical controls:</i>								
Agricultural suitability	yes	yes	yes	yes	yes	yes	yes	yes
Domesticated animals	yes	yes	yes	yes	yes	yes	yes	yes
Tropics	yes	yes	yes	yes	yes	yes	yes	yes
Political hierarchies	yes	yes	yes	yes	yes	yes	yes	yes
Economic complexity	yes	yes	yes	yes	yes	yes	yes	yes
<i>Contemporary controls:</i>								
In income, ln income ²	yes	yes	yes	yes	yes	yes	yes	yes
Communism indicator	yes	yes	yes	yes	yes	yes	yes	yes
Polity	no	no	no	no	yes	yes	no	no
Continent fixed effects	no	yes	no	yes	no	yes	no	yes
Observations	159	159	105	105	125	125	135 ^a	135 ^a
R-squared	0.41	0.43	0.15	0.21	0.28	0.32		

Notes: OLS estimates are reported with robust standard errors in brackets. The unit of observation is a country. ***, ** and * indicate significance at the 1, 5 and 10% levels. ^aThis is the average number of observations in the regressions for the three outcomes.

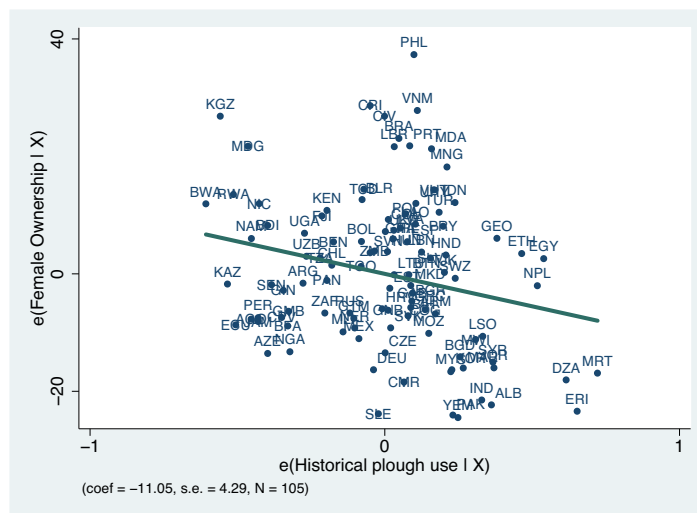
1, 3 and 5). From the figures it is clear that the coefficient for traditional plough use is not being influenced by a small number of countries. Further, the plots also show that the coefficient estimates are not only identified from broad differences across regions, but also from finer within-region variation. For example, we observe African countries in the Northwest corner (e.g., Rwanda, Madagascar) and in the Southeast corner (e.g., Eritrea, Mauritania, Ethiopia, etc). This is confirmed by the fact that the point estimates controlling for region fixed effects are essentially identical to the estimates without the fixed effects (comparing the odd numbered columns to the even numbered columns in Table 2).

Not only are the coefficient estimates statistically significant, but they are also economically meaningful. Based on the estimates from column 1, a one-standard-deviation increase in historical plough use (0.474) is associated with a reduction of female labor force participation (FLFP) of 7.82 percentage points (16.506×0.474), which is equal to 15.1% of the sample average for FLFP and 47% of its standard deviation. The impact on the share of firms with some female ownership (based on the column 3 estimates) is a reduction of 5.23 percentage points, which is 16% of the outcome's mean and 38% of its standard deviation. The reduction on the participation of women in politics (using the column 5 estimates) is 2.66 percentage points, which is 22% of the outcome's mean and 30% of its standard deviation.¹⁹

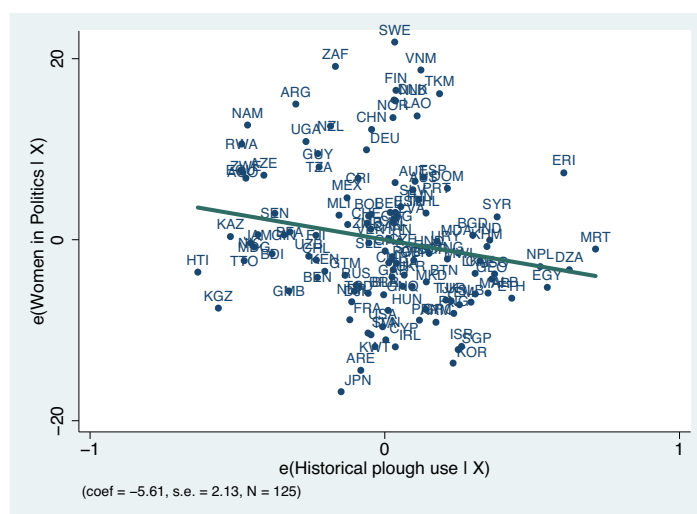
¹⁹See Appendix Table A1 for the means and standard deviations of the variables.



(a) Female labor force participation



(b) Female ownership or managing of firms



(c) Female participation in politics

Figure 5: Partial correlation plots for historical plough use, $Plough_c$.

Columns 7 and 8 report the estimated average effect size (AES) for the three dependent variables examined in columns 1–6. We computed the AES following Kling, Liebman, Katz and Sanbonmatsu (2004). Let β^k indicate the estimated plough coefficient for outcome variable k and σ^k the standard deviation of outcome k . Then, the average effect size is equal to $\frac{1}{K} \sum_{k=1}^K \frac{\beta^k}{\sigma^k}$, where K is the total number of outcome variables. To properly calculate the sample variance of the AES, the coefficients β^k are jointly estimated in a seemingly unrelated regression framework.²⁰ The AES estimates confirm the findings when examining the outcomes individually: historical plough use is associated with less female participation in activities outside of the home today. As well, the implied magnitudes are similar. According to the AES estimate, a one-standard-deviation increase in plough use is associated with an average decrease (for the three outcomes) of 0.40 standard deviations.

An alternative way to assess the magnitude of the estimates is to calculate the proportion of the total variation they explain. By this metric as well, historical plough use explains a sizable proportion of differences in gender roles across countries. When female labor force participation is the dependent variable (column 1 of Table 2), the inclusion of the historical plough use variable increases the R -squared by 0.086 (from 0.326 to 0.412). Therefore, traditional plough use accounts for 8.6% of the total variation in FLFP and 12.8% of the residual variation in FLFP unaccounted for by the control variables.²¹ For the share of firms with female ownership, traditional plough use accounts for 5% of the total variation and 6% of the residual variation. For the participation of women in politics, historical plough use explains 3% of the total variation and 4% of the residual variation.

Although we do not report them explicitly in here, we find that the estimated coefficients for the control variables are generally as expected.²² For example we find evidence of a U-shaped relationship between per capita income and female labor force participation, as well as the other outcomes. This is consistent with previous studies that also find this same non-monotonic relationship (Goldin, 1995). We also find that countries that experience a period of communism have higher rates of female labor force participation.

²⁰See Clingingsmith, Khwaja and Kremer (2009) for an alternative application and further details.

²¹This is calculated as: $(0.412 - 0.326) / (1 - 0.326) = 0.128$ or 12.8%.

²²The estimates for the controls are available upon request.

Robustness to alternative plough measures

Using either of the two methods for imputing missing language data yields estimates that are qualitatively identical to the estimates using our baseline variable. As reported in Appendix Table A2, the alternative measures yield nearly identical point estimates that are highly significant. As a final check to ensure that our findings are not being driven by measurement error, we omit 17 countries that have a significant proportion of missing language data. The countries include Australia, Argentina, Bolivia, Brazil, Chile, Colombia, Costa Rica, Ecuador, Guatemala, Honduras, Mexico, New Zealand, Nicaragua, Panama, Paraguay, Peru and Venezuela. The estimates, reported in Appendix Table A3, show that the impact of the plough remains robust to this check.

The persistence of female labor force participation

To this point, we have shown that historical plough use is associated with less female participation in agriculture historically and with less female participation in the labor force today. These two correlations suggest long-term persistence in female participation in activities outside of the home. As a check for this, we regress female labor force participation today on the measure of women's historical participation in agriculture constructed from the *Ethnographic Atlas*. The regression also controls for our full set of covariates from equation (2) above. The partial correlation plot, showing the relationship between historical female participation in agriculture and FLFP today, is shown in Figure 6. As is apparent from the figure, there is strong persistence over time. Female labor force participation today and female participation in agriculture in the past are very strongly correlated.

In light of existing studies, this persistence is perhaps surprising. In fact, Goldin and Sokoloff (1984) document that within the Northeastern United States, the low relative productivity of women and children in agriculture (and hence their low participation) spurred industrialization and their active participation in the manufacturing sector. In their setting, female labor force participation in agriculture was inversely related to participation in manufacturing, suggesting a lack of continuity of female labor force participation overtime as industrialization occurs. However, our results show that this example does not appear to be general. Instead, areas with low female participation in agriculture and plough use historically (because of the persistence of

patrilineal. All of the controls are constructed using the *Ethnographic Atlas* and in the same manner as the historical plough use and other ethnographic control variables.²³

Columns 1–3 of Table 3 report estimates of equation (2) with the additional controls included. For brevity, we only report estimates with female labor force participation as the dependent variable. (The estimates for the other outcomes variables are also robust to the inclusion of the additional controls.) The estimated impact of traditional plough use remains robust to the inclusion of the additional controls: the coefficient remains negative and statistically significant, and its magnitude changes little from the baseline value of 16.5.

It has been hypothesized that cultures with large-extended families typically have more hierarchical and less egalitarian structures, and since hierarchies tend to be dominated by men, this results in a subordinate status of women (Engels, 1902, Boserup, 1970, Barry, Bacon and Child, 1957).²⁴ We control for the potential impact of family structures by controlling for the proportion of a country’s ancestors that lived in nuclear or extended families.²⁵ Estimates with these additional controls are reported in column 4. The estimate for historical plough use remains robust.

It is possible that the status of women is also affected by the extent to which a society participates in certain non-agricultural activities, which are less suitable to female involvement. Likely candidates include the hunting and herding of large animals, which may require significant strength. We construct measures of the proportion of subsistence activities accounted for by hunting and by the herding of large animals.²⁶ As shown in column 5, the inclusion of these controls has little impact on the estimated effect of past plough use.

A potentially important determinant of differences in gender roles is religion. We check the robustness of our estimates by controlling for five variables that measure the proportion of a country’s population that is: Catholic, Protestant, other Christian, Muslim, and Hindu.²⁷ As shown in column 6, the estimated impact of the plough remains robust to controlling for current religion.

²³The controls are derived from variables v12, v28 and v75 from the *Ethnographic Atlas*. Full details of their construction are provided in the paper’s appendix.

²⁴A similar but different theory stresses competence rather than authority. In large families with many adults, a gender division of labor can more easily develop. In nuclear families with only a husband and a wife, it is more likely that either adult will need to substitute for the other. Therefore, the wife will be involved in activities ordinarily done by men (Whyte, 1978).

²⁵The information is taken from variable v8 of the *Ethnographic Atlas*.

²⁶The variables are constructed from variables v2 and v4 of the *Ethnographic Atlas*.

²⁷The data are taken from McCleary and Barro’s (2006) Religion Adherence dataset.

Table 3: Robustness of OLS estimates to alternative controls.

	Dependent variable: FLFP					
	(1)	(2)	(3)	(4)	(5)	(6)
Historical plough use	-14.734*** (4.856)	-16.814*** (3.769)	-17.073*** (3.568)	-14.570*** (3.568)	-15.417*** (3.692)	-12.041*** (3.967)
Practices intensive agriculture	yes					
Absence of private property		yes				
Patrilocal society, matrilineal society			yes			
Nuclear family, extended family				yes		
Proportion of subsistence from: hunting, herding					yes	
Prop of pop belonging to five major religions						yes
Baseline historical and contemporary controls	yes	yes	yes	yes	yes	yes
Observations	159	156	159	159	159	157
R-squared	0.41	0.41	0.42	0.46	0.43	0.57
	(7)	(8)	(9)	(10)	(11)	(12)
Historical plough use	-16.435*** (3.536)	-15.767*** (3.538)	-18.474*** (4.112)	-16.467*** (3.564)	-16.519*** (3.592)	-15.375*** (3.844)
Oil production per capita	yes					
Trade/GDP		yes				
Agric., manuf. and services share of GDP			yes			
Years of civil conflicts (1816-2007)				yes		
Years of interstate conflicts (1816-2007)					yes	
Fraction of European descent						yes
Baseline historical and contemporary controls	yes	yes	yes	yes	yes	yes
Observations	157	159	157	157	157	151
R-squared	0.43	0.42	0.42	0.41	0.41	0.42

Notes: OLS estimates are reported with robust standard errors in brackets. The unit of observation is a country. Each regression includes the full set of control variables (historical and contemporary). ***, ** and * indicate significance at the 1, 5 and 10% levels.

Ross (2008) challenges the view that religion – particularly Islam – has an important impact on gender roles. He argues instead for the importance of a country's endowment of oil reserves. According to his hypothesis, oil causes a country's domestic currency to strengthen, making exports less competitive and causing a decline in light manufacturing, a sector particularly well-suited for female employment. We account for this possibility by controlling for per capita oil production and the trade-to-GDP ratio, both measured in 2000. We also control for the economic structure of each country by including three variables that capture the share of GDP accounted for by agriculture, manufacturing and services.²⁸ Estimates with the additional controls are reported in columns 7–9. The impact of the plough remains robust.

Differences in gender roles could potentially be influenced by differences in the prevalence of warfare across societies. However, a priori, the direction of the effect is unclear (see Whyte, 1978). Involvement in warfare may cause societies to become more hierarchical and male dominated, suggesting a negative relationship between conflict and female work outside the home. On the other hand, being involved in warfare can generate a greater need for female involvement outside of the home. We control for the potential impacts of warfare, by calculating, for each country, the

²⁸The data are from the *World Development Indicators*, except for the oil production data which is from BP Oil (2006).

number of years since 1816 (the first year data are available) that the country was involved in either internal or interstate warfare.²⁹ The results, reported in columns 10 and 11, show that the impact of plough use is robust to controlling for warfare.³⁰

The last factor that we consider is motivated by the fact that country-level ancestral plough use is affected by historical European migration. European ancestry may have an effect on female labor force participation independent of historical plough use. We address this possibility by controlling for the fraction of each country's population in 2000 with ancestors that came from Europe.³¹ As reported in column 12, the estimates remain robust to controlling for European ancestry.³²

Overall, the estimated impact of the plough remains highly robust across the various specifications reported in Table 3. The coefficient is always negative and statistically significant, and the point estimates remain stable, ranging from -0.12 to -0.18 .

B. Individual-level estimates

We now turn to our specification that examines variation across individuals, linking them to a tradition of plough agriculture using the district they live in. The analysis relies on data from the *World Value Survey* (WVS), a compilation of national individual-level surveys on a wide variety of topics, including attitudes and preferences, as well as information on standard demographic characteristics, such as gender, age, education, labor market status, income, religion, etc.³³ Using the WVS we construct, for females, an indicator variable that equals one if she is in the labor force, which is defined as full-time, part-time or self-employment. Women are not in the labor force if they report being retired, a housewife or a student.³⁴

We also examine two measures of individuals' attitudes (male and female) about the appropriate role of women in society. The first measure is based on each respondent's view of the following

²⁹The data are from version 4 of the Correlates of War Database.

³⁰In Table 3, we control for the two types of warfare separately. Including both controls simultaneously does not alter the conclusion. In this case the coefficient is equal to -16.512 with a standard error of 3.578.

³¹The measure is taken from Nunn and Puga (2011), who calculate the variable using Putterman and Weil's (2010) *World Migration Matrix*.

³²An alternative, more brut-force strategy, is to omit all European and the neo-European countries (Australia, New Zealand, Canada, the US) from the sample. As reported in Appendix Table A4, the results also remain robust to this check.

³³Five waves of the WVS were carried out between 1981 and 2007. In our analysis, we use the four most recent waves of the survey, since the first wave does not contain information on the district in which the respondent lives. Because regional classifications often vary by wave, we use the wave with the most finely defined location data.

³⁴The results are qualitatively identical if we exclude retired women and students from the sample.

statement: “When jobs are scarce, men should have more right to a job than women”. The respondents are then asked to choose between agree, disagree, neither or don’t know. We omit observations for which the respondents answered ‘neither’ or ‘don’t know’, and code ‘disagree’ as 0 and ‘agree’ as 1.³⁵ Therefore, the constructed variable is increasing in the extent to which a respondent’s view is characterized by gender inequality.

We also consider a second variable derived from a survey question based on the following statement: “On the whole, men make better political leaders than women do”. Respondents are then asked to choose between ‘strongly disagree’, ‘disagree’, ‘agree’, ‘agree strongly’, or ‘don’t know’. We omit observations in which the respondent answered ‘don’t know’ and create a variable that takes on the value of 1 for ‘strongly disagree’, 2 for ‘disagree’, 3 for ‘agree’ and 4 for ‘agree strongly’. This variable, like the first, is increasing in the respondent’s view of gender inequality.

An appealing aspect of the two subjective belief-based variables is that they provide measures of the values that form the foundation of the objective outcome variables from the cross-country analysis. The first question reflects differences in individual beliefs about whether women should have equal access to jobs, which likely affect observed differences in female labor force participation rates across countries. The second question reflects values about the ability of women to take on roles of leadership and responsibility, which likely affects observed differences in female participation in politics and female firm ownership and management. Therefore, there is a close link between the objective measures from the country-level analysis and the subjective measures in the individual-level analysis.

Examining the three outcomes – female participation in the labor force, attitudes about female employment, and attitudes about female leadership – we estimate the following individual-level equation:

$$y_{i,d,c} = \alpha_{r(c)} + \beta \text{Plough}_d + \mathbf{X}_c^C \mathbf{\Gamma} + \mathbf{X}_d^H \mathbf{\Pi} + \mathbf{X}_i \mathbf{\Phi} + \varepsilon_{i,d,c} \quad (3)$$

where i denotes an individual, d denotes a district within a country c , and $r(c)$ denotes the continent of country c . Plough_d is our measure of traditional plough use among the ancestors of individuals living in district d . \mathbf{X}_c^C are the same contemporary country-level controls as in

³⁵We omit observations that respond ‘neither’ because it is ambiguous whether this represents an intermediate view or whether they have chosen not to answer the question or whether they do not know their answer. If we interpret this response as reflecting an intermediate position and code a variable that takes on the values 0, 1, and 2, then we obtain qualitatively identical results to what we report here.

Table 4: Individual-level OLS estimates.

	Dependent variables:			
	FLFP	When jobs are scarse	Men better political leaders	Average effect size (AES)
	(1)	(2)	(3)	(4)
Historical plough use	-0.214*** (0.034)	0.245*** (0.029)	0.397*** (0.075)	0.451*** (0.063)
Individual controls	yes	yes	yes	yes
Contemporary country controls	yes	yes	yes	yes
Historical district controls	yes	yes	yes	yes
Continent fixed effects	yes	yes	yes	yes
Observations	38,832	71,656	59,288	65,472
R-squared	0.19	0.21	0.18	

Notes: The table reports OLS estimates, with standard errors clustered at the country level. Individual controls are age, age squared, education, gender (for gender attitudes only), marital status, and income. Contemporary country controls include ln income, ln income squared and a communism indicator variable. Historical district controls include agricultural suitability, domesticated animals, tropical areas, political hierarchies, and economic complexity. The AES reported in column 4 is for the two subjective belief measures from columns 2 and 3. ***, ** and * indicate significance at the 1, 5 and 10% levels.

equation (2), and X_d^H includes the same historical ethnographic variables as in equation (2), but measured at the district level rather than the country level. X_i denotes current individual-level controls: age, age squared, as well as fixed effects for marital status, educational attainment, and income levels. The equation also includes continent fixed effects, denoted $\alpha_{r(c)}$. To be as conservative as possible, we cluster the standard errors at the country level.

Table 4 reports OLS estimates of equation (3). Consistent with the country-level estimates, we find a negative relationship between historical plough use and female labor force participation today, and a positive relationship between historical plough use and current attitudes reflecting gender inequality. In terms of the magnitude of the effects, they are similar to the cross-country estimates. A one-standard-deviation increase in the plough variable implies a reduction in female labor force participation of 0.09 or 9 percent (which is roughly equal to 16% of the sample average) and an increase in the female employment and female leadership gender attitudes variables of 0.10 and 0.21, respectively. These are 21% and 8% of the sample averages.

5. IV estimates

There is a number of concerns with the OLS estimates reported to this point. One is that locations that historically had attitudes prone to less equal gender roles may have been more likely to invent

or adopt the plough. This would bias our OLS estimates away from zero. It is also possible that locations that were economically more developed were more likely to have adopted the plough. Since these areas today are richer and more prone to attitudes about gender role equality, this would tend to bias our OLS towards zero.

Our first strategy to address these, and related, concerns is to control for observable characteristics. This is done in the previous section. In this section, we pursue an alternative strategy using instrumental variables. As instruments we exploit one determinant of historical plough use that has been emphasized in the anthropological literature: the type of crops grown in a particular location (Pryor, 1985).

The primary benefit of the plough is that it facilitates the cultivation of larger amounts of land over a shorter period of time. This capability is more advantageous for crops that require specific planting conditions that are only met during narrow windows of time or for crops that require larger tracts of land to cultivate a given amount of calories. The benefit of the plough is also reduced for crops grown in swampy, sloped, rocky, or shallow soils, all of which make the plough less efficient or impossible to use. Taking these factors into consideration, Pryor (1985) has classified crops into those whose cultivation benefits greatly from the adoption of the plough – he calls these plough positive crops – and those whose cultivation benefits less – called plough-negative crops. Plough-positive crops, which include wheat, teff, barley and rye, tend to be cultivated on large expanses of land (per calorie of output) that is flat, deep, with soil that are not too rocky or swampy, and have shorter growing seasons. Plough negative crops, which include sorghum, maize, millet, roots and tubers, and tree crops, tend to yield more calories per acre, have longer growing seasons, and can be cultivated on more marginal land (Pryor, 1985, p. 732).

Because the cultivation of plough-positive and plough-negative crops is an endogenous outcome, we do not use this as our instrument. Instead, we measure geo-climatic conditions that are unaffected by human actions, but which impact the suitability of a location for growing both types of crops. Our strategy uses two instruments. The first is a measure of the average suitability of the location of each observation's ancestors for cultivating plough positive cereal crops – wheat, barley and rye. The second is the same measure of ancestral suitability, but for cultivating millet and sorghum, which are plough-negative cereal crops. We intentionally choose to consider this set of crops because they are similar in many other dimensions except for the extent to which

they benefit from the use of the plough. Both sets of crops are cereals that have been cultivated in the Eastern Hemisphere since the Neolithic revolution (Mazoyer and Roudart, 2006, pp. 71–99). Both sets of crops require similar preparations for consumption, all being used for flour, porridge, bread or in beverages (Recklein, 1987). Both sets also produce similar yields and therefore neither clearly dominates the other in terms of the population it can support (Pryor, 1985, p. 732).

We obtain information on the suitability of a location for cultivating the plough-positive and plough-negative cereal crops from the FAO’s *Global Agro-Ecological Zones* (GAEZ) 2002 database (Fischer et al., 2002). The database reports suitability for the cultivation of numerous crops for 5 arc minutes by 5 arc-minute grid-cell globally. The measures are constructed from measures of the geo-climatic characteristics of a location: precipitation, frequency of wet days, mean temperature, daily temperature range, vapor pressure, cloud cover, sunshine, ground-frost frequency, wind speed, soil slope, and soil characteristics. The data are then combined with the specific growing requirements of crops to produce a measure of whether the crop can be grown in a location, and if so, how productively. It is important to note that the models of crop growth are based solely on previously established technical requirements and constraints for crop growth. The model’s parameters, and the final measures, are not affected by where crops are actually cultivated. As well, the final estimates are not simple functions of the geographic characteristics used, but are based on precise, highly non-linear crop-specific models of evapotranspiration, water-balance, temperature profiles, temperature growing periods, length of growing period, thermal regimes, etc. This last point is particularly important as it allows us to check the robustness of our IV results to controlling for important geo-climatic characteristics which may be different in plough-positive and plough-negative environments.³⁶

We construct the instruments by first identifying the land traditionally inhabited by each ethnic group in the *Ethnographic Atlas*. We continue to use all land within 200 kilometers of an ethnic group’s centroid and measure the amount of land within this area that can grow each of the cereal crops that comprise the instruments.³⁷ Let x_e^w , x_e^b , x_e^r , x_e^s , and x_e^m be the amount of land that can cultivate wheat, barley, rye, sorghum and millet, respectively. Further, let x^{all} be the amount of land that could grow any crop (i.e., the amount of arable land). The ethnicity-level measures of suitability for plough-positive crops is given by: $Area_e^{pos} = \frac{1}{3}(x_e^w + x_e^b + x_e^r)/x_e^{all}$. While ethnicity-

³⁶For a detailed discussion of the data and its use in a different application see Nunn and Qian (2011).

³⁷The GAEZ database measures suitability as a proportion of maximum attainable yield. We define locations that obtain at least 40% of the maximum yield as suitable. The results are robust to the use of other thresholds.

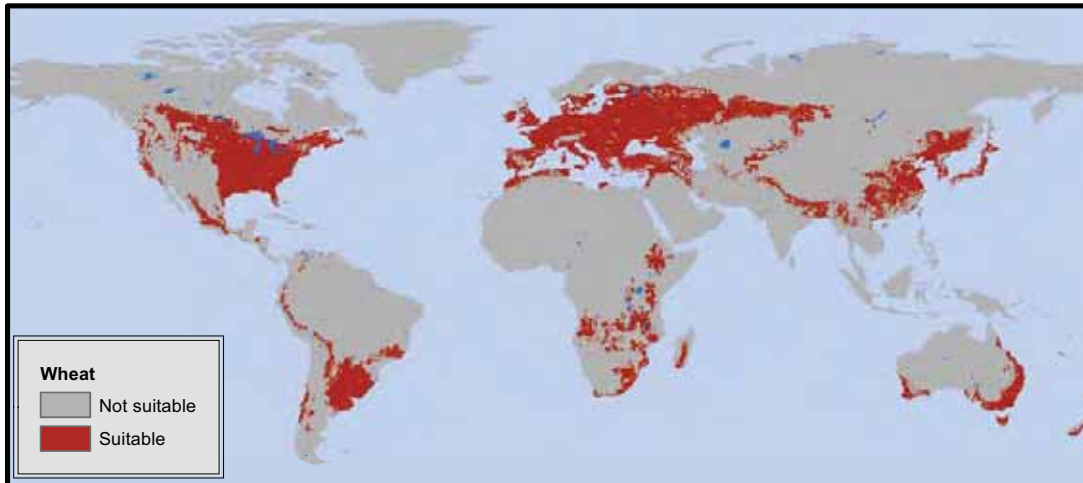
level measure of suitability for plough-negative crops is: $Area_e^{neg} = \frac{1}{2}(x_e^s + x_e^m)/x_e^{all}$. Intuitively, the instruments measure the average suitability for each type of crop, normalized by the overall suitability for cultivation in general.³⁸

Using the procedure explained in equation (1), we then construct district- and country-level averages of our plough-positive and plough-negative instruments. Intuitively, the instruments measure the proportion of the population with ancestors that had a climate that could grow plough-positive cereals (wheat, barley and rye) and the proportion that could grow plough negative cereals (sorghum and millet).

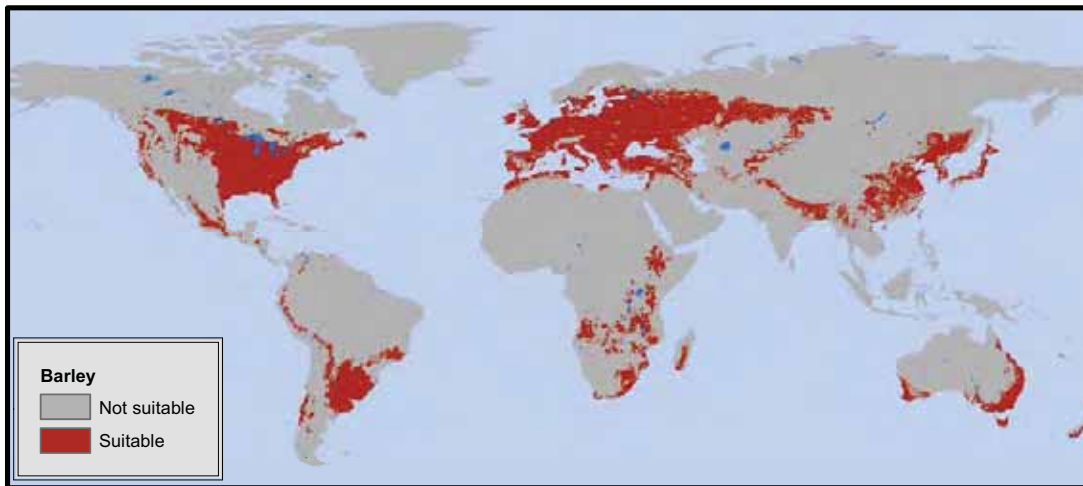
To provide the reader with a better sense of the instruments, Figure 7 shows the parts of the world that are classified as being suitable for the cultivation of the plough positive cereals – wheat, barley and rye – while Figure 8 shows suitability for the plough-negative cereals – millet and sorghum. A number of facts are apparent from the maps. First, there are many parts of the world that can grow plough-positive crops, but not plough-negative crops and vice versa. This provides an indication that the instruments may have variation independent from each other and therefore some predictive power. Second, relative to plough-positive crops, plough-negative crops appear to be relatively better suited for tropical and subtropical climates and plough-positive crops better suited for temperate climates. If these differences in climate affect gender attitudes today through channels other than past plough use, then the exclusion restriction will not be satisfied. Recall that the controls already include the proportion of land, historically inhabited by an ethnic group, that was tropical or subtropical. As well, in the analysis below we explicitly address this concern by exhaustively controlling for additional geographic characteristics that may be correlated with the suitability instruments.³⁹

³⁸Our procedure assumes that the GAEZ data provide an unbiased measure of historical suitability. In a different context, Nunn and Qian (2011) provide evidence of the validity of this assumption by showing that the GAEZ suitability measure for potatoes is highly correlated with historical potato production.

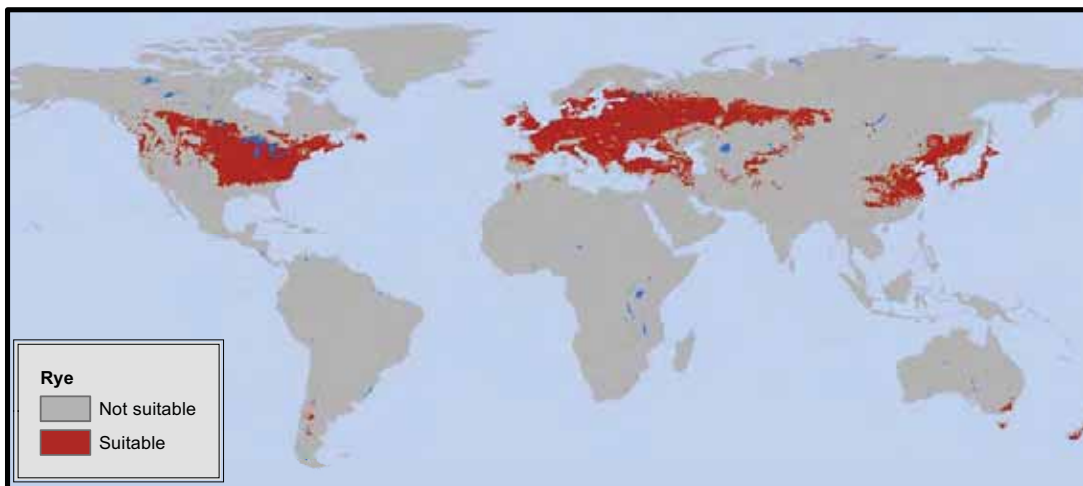
³⁹An important final point arises from the fact that the plough-positive and plough-negative cereals used in the construction of our instruments were all originally grown in the Eastern hemisphere and were not cultivated in the Americas until after 1500. This is not a concern to identification, but it is a fact that makes the first stage relationship weaker than it would be otherwise. For the large proportion of the population in the Americas whose ancestors are from the Eastern hemisphere, the instrument will provide predictive power. It is only for the indigenous populations of the Americas that the instrument will not affect plough adoption. This should be kept in mind when interpreting the IV estimates as a local average treatment effect (LATE). In other words, the estimates are an average effect among the ethnic groups whose plough adoption was affected by the geo-climatic suitability for growing the cereal crops.



(a) Wheat suitability

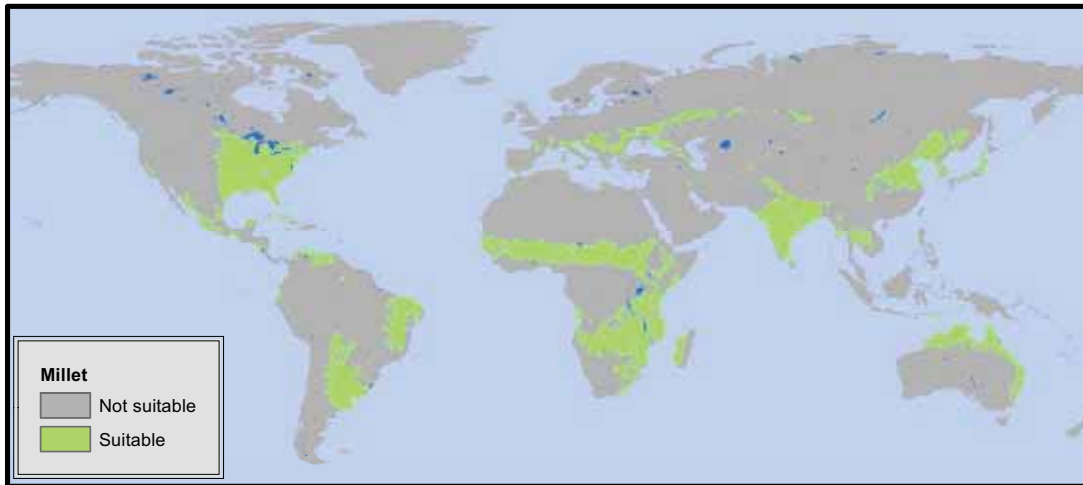


(b) Barley suitability

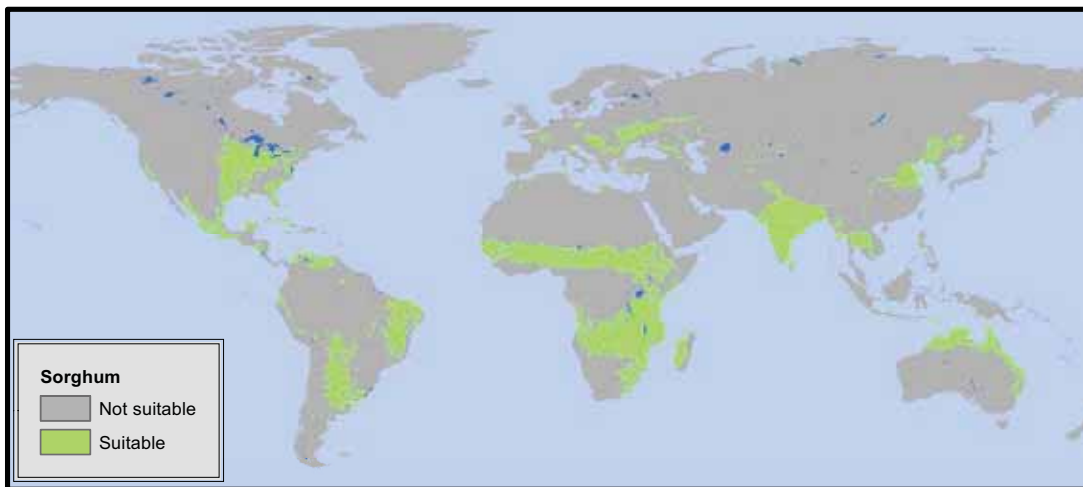


(c) Rye suitability

Figure 7: Maps displaying the global suitability of plough-positive crops, wheat, barley and rye.



(a) Millet suitability



(b) Sorghum suitability

Figure 8: Maps displaying the global suitability of plough-negative crops, millet and sorghum.

Table 5: Country level IV estimates.

Panel A. Second stage. Dependent variable:								
	Femal labor force participation		Share of firms with some female ownership		Females in politics	Average effect size (AES)		
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Historical plough use (2SLS)	-25.853** (10.051)	-26.423** (12.465)	-19.939* (11.932)	-26.274* (14.439)	-16.820* (9.243)	-23.089* (12.331)	-1.451*** (0.467)	-1.946*** (0.635)
Historical plough use (LIML)	-27.099	-28.25	-28.761	-43.12	-16.826	-23.16		
p-value	0.01	0.03	0.03	0.02	0.05	0.03		
CLR intervals	[-54.88, -7.02] [-66.24, -3.47] [-71.65, -2.70] [-142.60, -6.77] [-47.06, 0.29] [-75.07, -2.55]							
Historical controls	yes	yes	yes	yes	yes	yes	yes	yes
Contemporary controls	yes	yes	yes	yes	yes	yes	yes	yes
Continent FEs	no	yes	no	yes	no	yes	no	yes
Observations	157	157	104	104	124	124	133	133
R-squared	0.39	0.43	0.01	0.13	0.17	0.05		
Panel B. First stage. Dep var: Historical plough use								
Plough-positive environment	0.412*** (0.119)	0.377*** (0.101)	0.656*** (0.150)	0.561*** (0.143)	0.401*** (0.140)	0.340*** (0.117)		
Plough-negative environment	-0.120 (0.091)	-0.079 (0.075)	-0.017 (0.101)	0.001 (0.075)	-0.032 (0.103)	-0.026 (0.087)		
Equality of coefficients (p-value)	0.00	0.00	0.00	0.00	0.00	0.01		
F-stat (excluded instruments)	10.76	7.90	11.68	7.71	5.63	4.54		
Hausman test (p-value)	0.15	0.19	0.44	0.29	0.18	0.09		

Notes: IV estimates are reported with robust standard errors in brackets. The unit of observation is a country. Historical controls include agricultural suitability, domesticated animals, tropics, political hierarchies and economic complexity. Contemporary controls include ln income, ln income squared, a communism indicator, and polity (in columns 5 and 6 only). The number of observations for the AES estimates is the average number observations in the regressions for each outcome. ***, ** and * indicate significance at the 1, 5 and 10% levels.

A. Country-level estimates

Table 5 reports IV estimates of the specifications from Table 2. The first stage estimates are reported in the lower panel and the second stage estimates are in the top panel. The first stage estimates show that the historical suitability for the cultivation of plough-positive cereals is always positively correlated with the adoption of the plough, while suitability for the cultivation of plough-negative cereals is generally negatively correlated with plough use. In all specifications, the difference between the two coefficients is statistically significant. The F -test for joint significance of the two instruments is also reported in the table. The F -statistics range from about 5–11, suggesting that for some specifications there is a potential concern about weak instruments. For this reason we also report conditional likelihood ratio (CLR) confidence intervals and LIML estimates in addition to the regular 2SLS estimates.

The IV estimates, reported in panel A of the table, confirm the OLS estimates. Historical plough use is associated with less female labor force participation, less female ownership of firms, and less female participation in politics. The magnitude of the IV coefficients are consistently greater than the OLS estimates. This is potentially explained by selection arising from the endogeneity of plough adoption. All else equal, historically advanced societies were more likely

to adopt the plough. Further, they are more likely to also be advanced today with higher per capita incomes and more female participation in the labor market. Therefore, selection introduces a positive relationship between historical plough use and female labor force participation today, biasing the negative OLS estimates towards zero.

The IV estimates rest on the assumption that holding overall agricultural suitability constant, the specific type of cereal crop that a location could grow only impacts long-term gender attitudes through the past adoption of the plough. There are a number of potential concerns with this assumption. The first is that the difference between plough-positive and plough-negative environments may be correlated with geographic features that affect gender attitudes today through channels other than the plough. We check the robustness of our results to this concern by controlling for geographic characteristics that are potentially correlated with the suitability of the environment for plough-positive and plough-negative crops. Our controls include terrain slope, soil depth, average temperature and average precipitation of locations inhabited by each country's ancestors. Slope is measured as percent (i.e., rise over run). The soil depth control is the fraction of land that has no, few or slight soil depth constraints for cultivation. Average temperature is the average daily temperature (measured in degrees Celsius) between 1950 and 1959. Precipitation is the average monthly rainfall (in mm) over the same time period.

As reported in Table 6, the IV estimates remain robust to the inclusion of these geo-climatic characteristics. In columns 1–4, we control for each of the four measures individually. The point estimates remain positive and significant across each specification. In column 5, we control for all four controls simultaneously. Column 6 allows for the possibility of non-linear effects by also controlling for the square of each variable, while column 7 allows for interaction effects by also including all pairwise interactions of the controls. In each of the specifications, the points estimates remain robust, with magnitudes that are positive, significant and very similar in magnitude to the baseline estimate of 25.9.

We also check the robustness of the IV estimates to controlling for the additional covariates from Table 3. The results, which are reported in Table 7, show that the IV estimates, like the OLS estimates, remain robust to these additional controls.⁴⁰

⁴⁰The results are similarly robust for the other outcome variables as well.

Table 6: Robustness of IV estimates to additional geographic controls.

	Dependent variable: FLFP						
	(1)	(2)	(3)	(4)	(5)	(6)	(7)
Historical plough use	-33.573*** (11.777)	-25.539** (10.371)	-27.842*** (8.841)	-18.427** (9.068)	-25.089*** (9.300)	-26.076** (12.533)	-26.245** (10.696)
Terrain slope	yes				yes	yes	yes
Soil depth		yes			yes	yes	yes
Average temperature			yes		yes	yes	yes
Average precipitation				yes	yes	yes	yes
Quadratic terms						yes	
Linear interactions							yes
Baseline controls	yes	yes	yes	yes	yes	yes	yes
Observations	157	154	157	157	157	157	157
R-squared	0.29	0.38	0.38	0.38	0.44	0.44	0.44

Notes: OLS estimates are reported with robust standard errors in brackets. The unit of observation is a country. ***, ** and * indicate significance at the 1, 5 and 10% levels.

Table 7: Robustness of IV estimates to alternative controls.

	Second stage IV estimates: Dependent variable: FLFP					
	(1)	(2)	(3)	(4)	(5)	(6)
Historical plough use	-42.454* (22.737)	-25.563** (10.317)	-28.014*** (10.584)	-28.969*** (10.861)	-25.332*** (9.250)	-18.269** (9.003)
Intensity of agriculture	yes					
Absence of private property		yes				
Patrilocal and matrilineal societies			yes			
Nuclear and extended families				yes		
Proportion of subsistence from: hunting, herding					yes	
Fraction of major religious denomin.						yes
Baseline historical and contemporary controls	yes	yes	yes	yes	yes	yes
Observations	157	154	157	157	157	155
R-squared	0.27	0.41	0.38	0.38	0.40	0.55
	(7)	(8)	(9)	(10)	(11)	(12)
Historical plough use	-21.806* (11.188)	-25.974*** (10.014)	-25.606** (9.303)	-25.362** (9.930)	-25.578** (10.072)	-25.834** (10.602)
Fraction of European descent	yes					
Oil production		yes				
Trade/GDP			yes			
Agric., manuf. and services share of GDP				yes		
Years of civil conflicts (1816-2007)					yes	
Years of interstate conflicts (1816-2007)						yes
Baseline historical and contemporary controls	yes	yes	yes	yes	yes	yes
Observations	148	157	156	149	155	155
R-squared	0.44	0.40	0.40	0.43	0.40	0.40

Notes: IV estimates are reported with robust standard errors in brackets. The unit of observation is a country. Each regression includes the full set of control variables (historical and contemporary). ***, ** and * indicate significance at the 1, 5 and 10% levels.

Table 8: Individual-level IV estimates.

Panel A. Second stage. Dependent variable:				
	FLFP	When jobs are scarce	Men better political leaders	Average effect size (AES)
	(1)	(2)	(3)	(4)
Historical plough use	-0.585** (0.242)	0.576** (0.229)	1.383*** (0.362)	1.328*** (0.386)
Historical plough use (LIML)	-0.611	0.576	1.418	
p-value	0.00	0.00	0.00	
CLR intervals	[-0.700, -0.524]	[0.501, 0.652]	[1.308, 1.531]	
Individual controls	yes	yes	yes	yes
Contemporary country controls	yes	yes	yes	yes
Historical district controls	yes	yes	yes	yes
Continent FE	yes	yes	yes	yes
Observations	36,370	67,347	55,454	61,400
R-squared	0.15	0.17	0.11	
Panel B. First stage. Dependent variable: Historical plough use				
Plough-positive environment	0.273*** (0.073)	0.242*** (0.079)	0.338*** (0.088)	
Plough-negative environment	-0.075* (0.044)	-0.041 (0.040)	-0.094* (0.055)	
Equality of coefficients (p-value)	0.00	0.00	0.00	
F-stat (excluded instruments)	7.13	4.74	7.29	
Hausman test (p-value)	0.00	0.03	0.00	

Notes: The table reports IV estimates, with standard errors clustered at the country level. The instruments are plough-positive climate and plough-negative climate. Individual controls are age, age squared, education, gender (for gender attitudes only), marital status, and income. Current country controls include ln income, ln income squared and a communism indicator variable. Historical district controls include agricultural suitability, domesticated animals, tropical areas, political hierarchies, and economic complexity. The AES reported in column 4 is for the two subjective belief measures from columns 2 and 3. ***, ** and * indicate significance at the 1, 5 and 10% levels.

B. Individual-level estimates

Table 8 reports IV estimates of equation (3), which examines variation across individuals in the WVS. Consistent with the individual-level OLS estimates, the IV estimates also identify persistent impacts of historical plough use. We estimate a negative effect of past plough use on the participation of women in the labor force, and a positive effect on the prevalence of attitudes of gender inequality. Like the country-level estimates, we find that the IV estimates are larger than the OLS estimates.

6. Cultural transmission as a mechanism: Evidence from US immigrants

We now turn to an examination of the causal mechanisms underlying our results. Although our focus is on the evolution and persistence of cultural norms, it is possible that part of the long-term

effect of the plough may have arisen because it led to the development of institutions and markets which are less conducive to the participation of women in activities outside of the domestic sphere. Through this channel, the plough causes less female participation in market activities because it affects the costs and benefits of these activities, not because it affects individuals' beliefs about whether these are appropriate activities for women.

Our individual-level estimates, showing an impact of the plough on gender-role attitudes, provide evidence that the plough has impacted beliefs and values. However, these effects may have resulted from differences in external factors (e.g., institutions, policies, markets, etc), which in turn shape individual beliefs. To better isolate the causal impact of the plough on individual beliefs and values, we examine variation among second-generation immigrants – a group of individuals from diverse cultural backgrounds with different histories of ancestral plough use, but facing the same external environment, including markets, institutions and policies. Using data from the March Supplement of the *Current Population Survey* (CPS), we examine whether women with an ancestry of plough agriculture are less likely to be in the labor force today.⁴¹ We identify the ancestry of women in the sample using their parents' country-of-birth, measured using either the mother's country of birth, the father's country of birth, or both parents' country of birth (after restricting the sample to those with parents that have the same country of birth). One benefit of the different definitions of ancestry is that it allows us to examine whether cultural transmission is stronger from the father to the daughter, the mother to the daughter, or when both occur together.

Our analysis examines two samples of women, one that includes all women aged 15 to 64 and another that only includes married women in the same age group. When we examine married women only, we are able to also examine the impact of traditional plough use working through the husband's lineage. We identify a tradition of plough use among the husband's ancestor using either his father's country of birth, his mother's country of birth, or both.

We estimate the impact of the plough on immigrant populations using the following equation:

$$y_{i,s,c} = \alpha_s + \beta \text{Plough}_c + \mathbf{X}_c^C \Gamma + \mathbf{X}_c^H \Pi + \mathbf{X}_i \Phi + \varepsilon_{i,s,c} \quad (4)$$

where i denotes a second-generation woman currently living in state s , whose country of origin is country c . As in equation (2), Plough_c denotes the historical plough-use of those in country

⁴¹Starting in 1994, the CPS asks individuals about their country of origin and their parents' country of origin. We use all the years available since 1994.

c. α_s denote state fixed effects. The dependent variable, $y_{i,s,c}$, is an indicator variable that equals one if individual i is in the labor market. \mathbf{X}_c^C and \mathbf{X}_c^H denote the same vectors of covariates as in equation (2), which include current and historical ethnographic controls. \mathbf{X}_i indicates a vector of individual level controls, which includes dummies for education, a quadratic for age, real personal income, marital status, fixed effects for whether the person lives in a metropolitan or rural area, and fixed effects for the year of the survey. When we use the sample of married women we also include controls for characteristics of the husband: a quadratic of the husband's age, the husband's education, and his income. Because our variable of interest Plough_c only varies at the country of origin-level, we cluster all standard errors at this level.

The OLS estimates are reported in Table 9. Columns 1–3 report estimates using the full sample of women. Column 1 identifies the women's ancestry by her father's country of birth, while column 2 uses the mother's country of birth. In column 3, we restrict the sample to women whose parents were both born in the same country. For all three specifications, we estimate a negative relationship between a tradition of plough-use in the home country and the women's participation in the labor force. IV estimates of the same specifications, reported in columns 1–3 of Table 10, yield nearly identical results.

Columns 4–9 of Tables 9 and 10 report OLS and IV estimates using the sample of married women. Columns 4–6 reproduce the estimates of columns 1–3 for the sample of married women, when we identify ancestry using the wife's parents. We continue to find a negative impact of traditional plough use on female labor force participation, although we now find a smaller estimated impact when we identify ancestry using the father's country of birth. The impact is roughly 50% lower for both the OLS and IV estimates. This provides evidence that, at least for the subsample of married women, the transmission of values and beliefs (originating from the plough) is stronger from mothers to daughters than from fathers to daughters.

We next consider the possibility that a married woman's participation in the labor market may be influenced by her husband's beliefs and values, which were transmitted from his parents and their cultural background. Columns 7–9 of Tables 9 and 10, reproduce the estimates of columns 4–6 from the same tables, when we identify ancestry using the husband's parents rather than the wife's parents. The estimates provide evidence for an impact of a history of plough use among the husband's ancestors on the wife's participation in the labor market. Differentiating between the husband's parent's and their ancestry, we find a weaker estimate when we use the husband's

Table 9: Immigrant OLS regressions.

	Dependent variable: Labor force participation indicator								
	All women			Married women			Husband's ancestry		
	Women's ancestry			Women's ancestry			Husband's ancestry		
	Father's country	Mother's country	Parents same country	Father's country	Mother's country	Parents same country	Father's country	Mother's country	Parents same country
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
Historical plough use	-0.042*** (0.011)	-0.047*** (0.012)	-0.065*** (0.013)	-0.015 (0.021)	-0.043* (0.022)	-0.044* (0.024)	-0.027 (0.018)	-0.049*** (0.024)	-0.047*** (0.022)
Individual controls	yes	yes	yes	yes	yes	yes	yes	yes	yes
Husband controls	n/a	n/a	n/a	yes	yes	yes	yes	yes	yes
Historical country controls	yes	yes	yes	yes	yes	yes	yes	yes	yes
Contemporary country controls	yes	yes	yes	yes	yes	yes	yes	yes	yes
State fixed effects	yes	yes	yes	yes	yes	yes	yes	yes	yes
Observations	48,910	47,219	27,550	8,864	8,261	5,832	9,505	8,886	7,211
R-squared	0.37	0.38	0.40	0.37	0.37	0.39	0.37	0.37	0.38

Notes: OLS estimates are reported with standard errors clustered at the country level. An observation is a US second generation immigrant woman. Individual controls include age, age squared, education, marital status and income. Husband controls include husband's age, age squared, education and income. Historical country controls include the origin country's historical agricultural suitability, domestication of animals, tropics, political hierarchies and economic complexity. Contemporary country controls include ln income, ln income squared, and a communism indicator variable. All regressions also control for state, year and metropolitan area fixed effects. ***, ** and * indicate significance at the 1, 5 and 10% levels.

Table 10: Immigrant IV regressions.

	Panel A. Second Stage. Dep var: Labor force participation indicator											
	All women						Married women					
	Women's ancestry			Women's ancestry			Husband's ancestry			Husband's ancestry		
	Father's country	Mother's country	Parents same country	Father's country	Mother's country	Parents same country	Father's country	Mother's country	Parents same country	Father's country	Mother's country	Parents same country
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)
Historical plough use	-0.045*** (0.013)	-0.040** (0.016)	-0.058*** (0.019)	-0.026 (0.025)	-0.064** (0.029)	-0.060* (0.031)	-0.050** (0.025)	-0.079** (0.034)	-0.068** (0.030)			
Individual controls	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes
Husband controls	n/a	n/a	n/a	yes	yes	yes	yes	yes	yes	yes	yes	yes
Historical country controls	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes
Contemporary country controls	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes
State fixed effects	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes	yes
Observations	48,908	47,075	27,550	8,863	8,245	5,832	9,503	8,870	7,209			
R-squared	0.37	0.38	0.40	0.37	0.37	0.39	0.37	0.37	0.38			
	Panel B. First stage. Dep var: Historical plough use											
Plough-positive environment	1.195*** (0.168)	1.160*** (0.176)	1.297*** (0.176)	1.211*** (0.164)	1.223*** (0.170)	1.27*** (0.173)	1.194*** (0.152)	1.205*** (0.150)	1.241*** (0.149)			
Plough-negative environment	-0.370* (0.212)	-0.360 (0.222)	-0.421* (0.227)	-0.402** (0.188)	-0.372** (0.187)	-0.444** (0.186)	-0.361** (0.186)	-0.348** (0.180)	-0.394** (0.187)			
Equality of coefficients (p-value)	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00			
F-stat (excluded instruments)	60.02	48.16	65.80	64.45	54.43	61.82	66.47	71.19	85.44			
Hausman test (p-value)	0.73	0.50	0.54	0.64	0.34	0.46	0.21	0.19	0.22			

Notes: IV estimates are reported with standard errors clustered at the country level. An observation is a second generation US immigrant woman. Individual controls include age, age squared, education, marital status and income. Husband controls include husband's age, age squared, education and income. Historical country controls include the origin country's historical agricultural suitability, domestication of animals, tropics, political hierarchies and economic complexity. Contemporary country controls include ln income, ln income squared, and a communism indicator variable. All regressions also control for state, year and metropolitan area fixed effects. ***, ** and * indicate significance at the 1, 5 and 10% levels.

father's country of birth than the husband's mother's country of birth. Like the case for the women's parents, these estimates identify weaker transmission from the father than from the mother. In other words, the transmission of beliefs about gender appears stronger from mothers (than fathers), whether it is from the wife's mother to the wife or the husband's mother to the husband.

It is also interesting that, in every specification, the estimated impact of the plough transmitted through the husband's parents is stronger than the impact of the plough transmitted through the wife's parents (i.e., compare columns 4–6 with 7–8). This suggests that the beliefs of the husband are at least as important, if not more important, than the beliefs of the wife in determining whether she enters the labor force. This result is consistent with the previous findings of Fernandez and Fogli (2009).

Comparing the magnitude of the immigrant estimates with the country-level estimates allows us to glean some evidence of the relative importance of cultural transmission, relative to other channels, in explaining the relationship between historical plough use and female labor force participation. The impact of the plough on US second-generation immigrant women is much smaller than the estimated impact from the country-level regressions. Consider the estimated impact on female labor force participation from a one-unit change in historical plough use (i.e., an increase zero to one). Using the country-level OLS estimates (column 1 of Table 2), this is associated with an increase in FLFP by 16 percentage points. The individual-level estimates (column 1 of Table 4) suggest impacts of a similar magnitude: 21 percentage points. However, the magnitudes based on the immigrant-level estimates are much smaller in magnitude, ranging from 4 to 6 percentage points (columns 1–9 of Table 9). Although one must interpret these findings with caution, the relative magnitudes suggest that internal norms (identified in the immigrant regressions) account for approximately a quarter to a third of the overall impact of the plough on gender roles.

7. Conclusions

Social anthropologists have long-considered the use of shifting hoe cultivation vs. plough cultivation as an important determinant of the evolution and persistence of traditional gender roles and norms. We formally test this hypothesis by combining ethnographic data on traditional plough

use, measured at the ethnicity level, with contemporary data measuring gender attitudes and female participation outside of the domestic sphere.

Our findings provide evidence that current differences in gender attitudes and female behavior have indeed been shaped by historical differences in agricultural systems. Specifically, we have shown that individuals, ethnicities and countries whose ancestors used the plough have beliefs that exhibit greater gender inequality today and have less female participation in non-domestic activities, like market employment, entrepreneurship, and politics. In an effort to identify a channel of cultural persistence, we examined variation across second-generation female immigrants born and living in the US, but from different cultural backgrounds. We find that even among this group of individuals, who face the same labor market, institutions, and policies, a history of plough use is associated with less female labor force participation.

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